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## Applying a dynamic ARDL approach to the Environmental Phillips Curve (EPC) hypothesis amid monetary, fiscal, and trade policy uncertainty in the USA

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1 **Applying a dynamic ARDL approach to the Environmental Phillips Curve**  
2 **(EPC) hypothesis amid monetary, fiscal, and trade policy uncertainty in the**  
3 **USA**

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**27 Abstract**

28 It is well known that unemployment and environmental degradation are two critical issues across  
29 the globe. However, there is extended dearth of the literature that explores the nexus between  
30 unemployment and environmental degradation. Kashem and Rahman (2020) put forward the  
31 Environmental Phillips Curve (EPC) hypothesis, which depicts a negative relationship between  
32 unemployment and environmental degradation. This study further explores the validity of the EPC  
33 hypothesis in the case of the US. It also investigates the impact of monetary policy uncertainty  
34 (MU), fiscal policy uncertainty (FU), and trade policy uncertainty (TU) on carbon dioxide  
35 emissions. To this end, the analysis employs the novel methodology of dynamic ARDL model.  
36 The results document that EPC does not hold in the short-run, but it does in the long-run.  
37 Furthermore, both in the short- and long-run, MU escalates CO<sub>2</sub> emissions, while FU plunges  
38 emissions in both the short- and long-run. Finally, TU does not alter the level of CO<sub>2</sub> emissions.

39 **Keywords:** Environmental Phillips Curve; monetary policy uncertainty; fiscal policy uncertainty;  
40 trade policy uncertainty; CO<sub>2</sub> emissions; Dynamic ARDL model

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## 50 **Introduction**

51 Global warming, climate change, and environmental degradation are repeatedly cited as the  
52 emerging concerns for humanity nowadays. These factors are mainly responsible for the rise in  
53 global temperature and extreme weather conditions. Moreover, food insecurity has also been  
54 escalating due to climate change and environmental degradation. In addition, these environmental  
55 issues have further detrimental effects on human health (Warner et al., 2010; Azam, 2016; Kompas  
56 et al., 2018; Gorus and Aslan, 2019; Khavarian et al., 2019), while climate change also impacts  
57 production and consumption activities (Alshehry and Belloumi, 2015; Shahbaz et al., 2017;  
58 Afionis et al., 2017).

59 One of the prime reasons behind climate change, global warming, and environmental change  
60 is greenhouse gases (GHGs) emissions, with the carbon dioxide emissions being the most critical  
61 among all GHGs. It is worth noting that the share of CO<sub>2</sub> emissions in total GHGs comes as much  
62 high as 80% (Gill et al., 2018; Sarkodie and Strezov, 2019), implying that controlling CO<sub>2</sub>  
63 emissions is imperative in order to curb overall GHGs emissions. International organizations, with  
64 the support of developed and developing countries, have been trying to mitigate the levels of CO<sub>2</sub>  
65 emissions through several initiatives (Kyoto protocol and Paris agreement); however, the levels of  
66 CO<sub>2</sub> emissions have been gradually rising across the globe. This concern calls for further actions,  
67 and especially propels to probe the key determinants of CO<sub>2</sub> emissions.

68 The literature on environmental economics reports that both economic growth and energy  
69 consumption are the prime reasons for high CO<sub>2</sub> emissions (Belke et al., 2011; Zhu et al., 2016).  
70 Economic growth, aided by non-renewable energy consumption, deteriorates the environmental  
71 quality (Zafar et al., 2020; Destek and Sinha, 2020; Mahalik et al., 2021). More specially, energy  
72 consumption exploits crude oil, coal, and natural gas' resources, which emit enormous CO<sub>2</sub>

73 emissions. Likewise, economic growth allows consumers to consume pollution-intensive goods,  
74 which eventually lead to higher levels of CO<sub>2</sub> emissions. Hence, there exists an economy-  
75 environment dilemma, according to which, mitigating CO<sub>2</sub> emissions might affect economic  
76 growth. Therefore, it is imperative to explore other determinants of CO<sub>2</sub> emissions, so that they  
77 could be reduced without affecting economic growth.

78 The current literature put forwards several drivers of CO<sub>2</sub> emissions, such as financial  
79 development (Shahbaz et al., 2013a, b; Abbasi and Riaz, 2016; Dogan and Turkekul, 2016; Bekhet  
80 et al., 2017; Shoaib et al., 2020), trade (Halicioglu, 2009; Shahbaz et al., 2013a; Chen et al., 2019;  
81 Haug and Ucal, 2019), natural resources (Bekun et al., 2019; Danish et al., 2019; Khan et al.,  
82 2020), urbanization (Zhu et al., 2012; Sadorsky, 2014; Shahbaz et al., 2016; Ali et al., 2019),  
83 population (Dietz and Rosa, 1997; Begum et al., 2015; Dogan and Kan, 2018; Hashmi and Alam,  
84 2019), energy prices (Zhang and Lin, 2012; Al-Mulali et al., 2013; Joo et al., 2015; Wang et al.,  
85 2016; Anser, 2019), energy efficiency (Khan et al., 2019; Nathaniel and Iheonu, 2019; Wolde and  
86 Weldemeskel, 2020), institutional quality (Abid, 2016; Bhattacharya et al., 2017), corruption  
87 (Sekrafi and Sghaier, 2018; Wang et al., 2018; Arminen and Menegaki, 2019), terrorism (Bildirici,  
88 2020), globalization (Zaidi et al., 2019), and geopolitical risks (Adams et al., 2020). However,  
89 there exists a dearth of the literature on the nexus between unemployment and CO<sub>2</sub> emissions.

90 In addition, through the public and fiscal policy adjustment, such as the application of  
91 environmental taxes, has been found to potentially generate both environmental quality and  
92 employment or unemployment inferences. For instance, the double dividend hypothesis (DDH)  
93 though the application of environmental taxes, could improve or worsen environmental quality,  
94 while also increasing unemployment, depending on the understudied case (Carraro, et al., 1996;  
95 Degirmenci and Aydin, 2021). Accordingly, Schneider (1997) implies that the involuntary

96 unemployment and environmental quality aspects of an ecological tax policy are presentable  
97 through the efficiency wage model. However, the dimension of unemployment-environment  
98 nexus, which has been sparsely presented in the literature, exhibits a close semblance or cut out  
99 from the aforementioned DDH. Illustratively, Forstater (2003) informs that full employment, along  
100 with desirable environment quality, is attainable through a public service employment program  
101 based on functional finance. In essence, attaining a trade-off relationship between unemployment  
102 and environmental degradation could essentially require special attention from various  
103 stakeholders (i.e., policy makers and government officials, among others). Thus, attaining a low  
104 unemployment rate without compromising the environment necessitates the implementation of  
105 desirable policies or reforms.

106       Recently, Kashem and Rahman (2020) put forward the Environmental Phillips Curve (EPC)  
107 hypothesis, according to which there exists a negative relationship between unemployment and  
108 environmental quality. Anser et al. (2021a) also report the validity of this hypothesis in the case  
109 of the BRICST countries. The proponents of EPC claim that high unemployment rates mitigate  
110 production and, hence, plunge CO<sub>2</sub> emissions. In contrast, increases in unemployment plunge  
111 consumers' income. As a result, the willingness to pay for improved environmental quality  
112 declines and, hence, carbon emissions are expected to increase. Furthermore, unemployment may  
113 increase or decrease environmental quality, and it is inevitable to empirically probe the  
114 unemployment-environment nexus to device policies for sustainable development.

115       Moreover, Kashem and Rahman (2020) employ economic growth and trade as determinants  
116 of carbon emissions, while modeling the unemployment-environment nexus. However, to prevent  
117 the empirical/econometric model from misspecification and to test the validity of this nexus in the  
118 presence of other key determinants, additional important variables (e.g., industrial production,

119 energy consumption, and economic policy uncertainty) can be also incorporated. Recently,  
120 economic policy uncertainty (EPU) has emerged as one of the key influencing factors of  
121 environmental degradation (Jiang et al., 2019a). EPU has both economic and environmental  
122 effects, by altering the behavior of both consumers and producers, as well as that of environmental  
123 quality. Several studies have highlighted many channels capable of explaining the theoretical  
124 relationship between EPU and CO<sub>2</sub> emissions (Wang et al., 2020; Yu et al., 2021). In addition, a  
125 strand of the literature reports that EPU leads to higher CO<sub>2</sub> emissions (Danish et al., 2020; Anser  
126 et al., 2021b), whereas another group highlight that EPU impedes CO<sub>2</sub> emissions (Syed and Bouri,  
127 2021; Yu et al., 2021). By contrast, a different strand of the empirical literature expounds the  
128 absence of any significant link between EPU and CO<sub>2</sub> emissions (Abbasi and Adedoyin, 2021).  
129 While devising policies related to the environmental impacts of EPU as implied above, the  
130 evidence has posited contrasting and inconsistent results. Hence, further investigation on the effect  
131 of EPU on the environment within the newly proposed EPC framework is needed to reshape the  
132 current policies on how to achieve sustainable development goals.

133       Based on the above milieu, the objective of this study is to scrutinize the validity of the EPC  
134 hypothesis, while probing the impact of EPU on CO<sub>2</sub> emissions in the US. The motivation to  
135 choose the US as the case study here is based on the fact that this country is the largest economy  
136 of the world, as well as the second largest carbon emitter on a worldwide basis. Additionally, the  
137 categorical components of EPU that include the fiscal, monetary, and trade policy were employed,  
138 while considering the test for the validity of the EPC hypothesis. It is worth mentioning that there  
139 exist a few limitations of the study by Kashem and Rahman (2020). First, their study uses static  
140 analysis and does not report the dynamic relationship between unemployment and environmental  
141 quality. Second, their study omits several key drivers of carbon emissions, such as uncertainty

142 related to economic policies. In addition, US economic policies have turned to be substantially  
143 uncertain due to several external and internal issues (such as, the US-China trade war, the 2008  
144 global financial crisis, and the twin deficit crisis). This study contributes to the literature in a sense  
145 that the findings are expected to help policy makers to devise reforms and policies to curb CO<sub>2</sub>  
146 emissions, which eventually lead to the achievement of sustainable development goals.

147       The contribution of this study to the existing literature is threefold: i) It explores the validity  
148 of the novel EPC hypothesis in the case of the US, ii) prior studies on the nexus of EPU and the  
149 environment considered only aggregate measures of economic policy uncertainty except for Alola  
150 (2019 a & b) that explored the only monetary uncertainty (MU) and trade uncertainty (TU) without  
151 the economic uncertainty. The economic intuition behind exploring the disaggregate EPU is that  
152 any shocks in EPU are in fact shocks in its components. Moreover, the recent changes in trade and  
153 economic policies of the United States arising from the country's foreign policy redirection  
154 especially with China and other major economies are enough reasons to consider the disaggregate  
155 EPU. Thus, it might be possible that the impact of these three types of uncertainty on CO<sub>2</sub>  
156 emissions is heterogeneous, in a sense that only one type of policy uncertainty surges CO<sub>2</sub>  
157 emissions, while the others plunge them. Therefore, we fill this line of research and examine the  
158 effect of MU, TU, and FU within the EPC framework, and iii) this study employs the novel  
159 methodology of dynamic ARDL simulations for robust and efficient findings. Compare to previous  
160 ARDL modelling methods (e.g., standard ARDL, bootstrap ARDL, NARDL, QARDL) this  
161 approach allows to graphically explore the effect of shocks in the regressors on the predicted value  
162 of the dependent variable, which can be useful for certain future policy actions. Furthermore, this  
163 approach renders reliable, efficient, and robust results even with small sample datasets, since this  
164 approach is a simulation-based algorithm.



## 165 **Theoretical linkages**

166 This section presents the theoretical linkage among the considered variables of this study. In  
167 particular, it proposes two channels that can theoretically connect unemployment with  
168 environmental quality. According to the ‘growth channel’, unemployment impedes economic  
169 growth, which on the other hand mitigates energy consumption. As a result, carbon emissions are  
170 expected to decrease. In contrast, the ‘preference channel’ argues that unemployment diminishes  
171 consumers’ income, which does not allow individuals/households to express preferences in favor  
172 of improved environmental quality through expensive environmentally friendly goods.

173 On the other hand, there are several channels through which policy uncertainty affects CO<sub>2</sub>  
174 emissions. Jiang et al. (2019a) note that uncertainty about economic policy affects CO<sub>2</sub> emissions  
175 through the ‘direct policy adjustment’ effect, as well as the ‘indirect economic demand’ effect.  
176 The former illustrates that increased levels of EPU divert the attention of policymakers from  
177 environmental protection to economic stability, which renders opportunities for producers to use  
178 conventional (i.e., non-renewable energy) energy for production, thus, producing higher CO<sub>2</sub>  
179 emissions. The latter highlights that uncertainty regarding economic policy amends the decision-  
180 making process of economic agents, and thus, affects energy consumption, which eventually  
181 affects CO<sub>2</sub> emissions. In addition, Wang et al. (2020) put forward two other effects: i) the  
182 consumption effect, and ii) the investment effect. The former notes that EPU decreases the use of  
183 energy and pollution-intensive goods, which in turn leads to reduced CO<sub>2</sub> emissions. By contrast,  
184 the latter highlights that EPU has adverse effects on investments in R&D and renewable energy,  
185 which eventually upsurge CO<sub>2</sub> emissions.

186 Parallel to this, Yu et al. (2021) also propose three channels that bridge economic policies  
187 uncertainty with CO<sub>2</sub> emissions. These channels include: the innovations channel, the share of the

188 fossil fuel energy channel, and the energy intensity channel. The first one supports that EPU leads  
189 to less innovations, and hence to higher CO<sub>2</sub> emissions. The second one explains that EPU  
190 increases the share of non-renewable energy in the energy mix, and thus, CO<sub>2</sub> emissions are  
191 expected to increase. Finally, the third channel expounds that EPU raises energy intensity, which  
192 escalates CO<sub>2</sub> emissions.

### 193 **Literature review**

194 This section is segmented into two subsections. The first one highlights the influencing factors of  
195 CO<sub>2</sub> emissions, while a number of studies on the nexus between EPU and CO<sub>2</sub> emissions are  
196 presented in the second part.

### 197 **Determinants of CO<sub>2</sub> emissions**

198 As climate change and global warming are increasing concerns across the world, a substantial  
199 number of researchers have analyzed them along with different influential factors impelling carbon  
200 emissions (Richmond and Kaufmann, 2006; Katircioğlu and Taşpınar, 2017; Mutascu, 2018; Jiang  
201 et al., 2019b). In the economy-environment nexus, the Environmental Kuznets Curve (EKC)  
202 hypothesis has been a prime conjecture (Dogan and Turkekul, 2016; Pata, 2018; Işık et al., 2019),  
203 which implies the presence of an inverted U-shaped relationship between income and  
204 environmental degradation. Researchers have been investigating the validity of the EKC  
205 hypothesis over the last decades and have generated mixed and contrasting results. One group  
206 report that an inverted U-shaped relationship between income and environment does exist (Tang  
207 and Tan, 2015; Bilgili et al., 2016; Kacprzyk and Kuchta, 2020), while the other group claim that  
208 the presence of a N-shaped relationship is valid (Lee and Oh, 2015; Allard et al., 2018). Several  
209 other studies expound the U-shaped and roughly M-shaped relationship between income and  
210 environment quality (Sinha et al., 2017; Minlah and Zhang, 2021). It is worth mentioning that

211 models and methods, time period, countries, and the choice of control variables are mainly  
212 responsible for the mixed findings in the context of the EKC hypothesis (Heidari et al., 2015; Jamel  
213 and Maktouf, 2017; Pata, 2018).

214 Similarly, there are a few other studies that link (un)employment with the environment. More  
215 specifically, Witzke and Urfei (2001) examine the determinants of the willingness to pay for  
216 environmental issues, and they find the employment status is explicitly considered as one of those  
217 determinants. Likewise, Veisten et al. (2004) report that unemployment impedes the willingness  
218 to pay for high environmental quality. In contrast, there exists some empirical evidence which  
219 notes that the employment status and willingness to pay for environmental issues do not have any  
220 relationship between them (Torgler and García-Valiñas, 2007; Ferreira and Moro, 2013; De Silva  
221 and Pownall, 2014). Recently, Kashem and Rahman (2020) put forward the Environmental Phillips  
222 Curve (EPC) hypothesis, i.e., the presence of a negative relationship between unemployment and  
223 environmental quality. Additionally, Joshua and Alola (2020) examine the role of employment  
224 within the pollution haven hypothesis for the case of South Africa. They provide evidence that  
225 employment leads to high carbon emissions. Similarly, Gyamfi et al. (2020) use the EKC  
226 framework to investigate the relationship between employment and environment. The findings  
227 from this study document that rises in employment contribute to high carbon emissions. Next,  
228 Anser et al. (2021a) support the validity of EPC for the case of BRICST countries, and also report  
229 that economic growth and energy consumption escalate environmental degradation. In contrast,  
230 our study probes the impact of uncertainty related to economic policies within the EPC framework,  
231 whilst employing the novel dynamic ARDL simulations approach. In other words, our study  
232 extends the EPC literature in certain dimensions.

233 Parallel to this, energy consumption is often cited as one of the eminent drivers of CO<sub>2</sub>  
234 emissions (Saboori et al., 2014). The use of crude oil, natural gas, and coal emits high levels of  
235 CO<sub>2</sub> emissions (Destek and Sinha, 2020; Haug and Ucal, 2019). Several works also highlight the  
236 direction of causality between energy and environment (Zhang and Lin, 2012; Nathaniel and  
237 Iheonu, 2019). Moreover, one strand of the literature disaggregates energy into renewable and non-  
238 renewable energy and notes that these two energy sources have a heterogeneous impact on CO<sub>2</sub>  
239 emissions (Sadorsky, 2014). Likewise, energy efficiency (i.e., the productivity of energy  
240 consumption) plunges CO<sub>2</sub> emissions, since the same amount of energy can produce higher output  
241 (Afionis et al., 2017). Higher energy prices also can reduce the demand for energy, which  
242 eventually mitigates CO<sub>2</sub> emissions (Joo et al., 2014; Dogan and Turkekul, 2016).

243 Foreign direct investment (FDI) can either upsurge or impede CO<sub>2</sub> emissions. According to  
244 the pollution haven hypothesis, FDI could bring in environment unfriendly technologies. As a  
245 result, the levels of CO<sub>2</sub> emissions can significantly increase (Khavarian et al., 2019; Destek and  
246 Sinha, 2020). In contrast, the pollution Halo hypothesis notes that FDI encourages environment  
247 friendly technologies, and ultimately reduces CO<sub>2</sub> emissions (Belke et al., 2011; Jiang et al.,  
248 2019b). The environmental impact of trade is also unclear, because a strand of the literature argues  
249 that trade escalates environmental quality, while others reports that the opposite holds (Chen et al.,  
250 2019). More specifically, the trade-environment nexus depends on the nature of goods and services  
251 traded, as well as on the direction of trade (Halicioglu, 2009; Zhao et al., 2018).

252 Moreover, several studies report that natural resources also affect CO<sub>2</sub> emissions (Bekun et  
253 al., 2019; Danish et al., 2019). They note that urbanization escalates energy consumption, and  
254 hence, increases CO<sub>2</sub> emissions (Sadorsky, 2014; Ali et al., 2019). By contrast, others document  
255 that urbanization calls for improved environmental quality, which propels policy makers to initiate

256 strict environmental measures, leading to reduced CO<sub>2</sub> emissions (Richmond and Kaufmann,  
257 2006; Zhu et al., 2016; Bekhet et al., 2017). Likewise, tourism may also increase energy  
258 consumption due to the demand for infrastructure and transportation, which surges CO<sub>2</sub> emissions  
259 (Dietz and Rosa, 1997). Next, political, social, and economic globalization can also affect  
260 consumption and production decisions, and ultimately hit CO<sub>2</sub> emissions (Bilgili et al., 2016; Zaidi  
261 et al., 2019). The empirical literature also reports that political instability affects various economic  
262 decisions, and in turn, CO<sub>2</sub> emissions (Wang et al., 2018; Mahalik et al., 2021). Additionally,  
263 corruption, terrorism, and militarization can determine the levels of CO<sub>2</sub> emissions (Bildirici,  
264 2020), while monetary, fiscal, and trade policies can also have direct, as well as indirect, impacts  
265 on CO<sub>2</sub> emissions (Halicioglu, 2009; Dogan and Turkekul, 2016). Finally, a few other research  
266 outlets also show that there exists an asymmetric impact of economic policies on CO<sub>2</sub> emissions  
267 (Danish et al., 2019).

268 The expansion of R&D investment, innovations, and technological advancements could  
269 improve energy efficiency, with these factors being able to put forward new methods to utilize  
270 renewable energy. As a result, CO<sub>2</sub> emissions are expected to get significantly plunged (Garrone  
271 and Grilli, 2010; Zhang and Zhang, 2018). Furthermore, financial development can also promote  
272 green investments, which reduce CO<sub>2</sub> emissions. By contrast, there exist a few empirical studies  
273 which report that financial development upsurges energy consumption and economic growth,  
274 therefore, escalating the levels of CO<sub>2</sub> emissions (Shahbaz et al., 2013a; Bekhet et al., 2017; Shoaib  
275 et al., 2020).

## 276 **Economic policy uncertainty and environment**

277 There are several studies that explore that relationship between EPU and the environment. EPU  
278 may disturb the decision making of economic agents, which affects CO<sub>2</sub> emissions (Jiang et al.,

279 2019a). EPU can also affect CO<sub>2</sub> emissions through the energy efficiency channel (Danish et al.,  
280 2020). One strand of the literature on the EPU-environment nexus notes the positive relationship  
281 between EPU and the environment, while the other highlights the opposite. For instance, Adams  
282 et al. (2020) report that EPU escalates CO<sub>2</sub> emissions in the top resource rich economies. Similarly,  
283 Anser et al. (2021b) employ the STIRPAT model and note that EPU surges CO<sub>2</sub> emissions in the  
284 top ten carbon emitter countries. Likewise, using the world uncertainty index as a proxy for EPU,  
285 Wang et al. (2020) expound that EPU increases CO<sub>2</sub> emissions in the US. In addition, Adedoyin  
286 and Zakari (2020) employ the ARDL approach and incorporate these findings in the case of the UK.

287 Syed and Bouri (2021) use the bootstrap ARDL approach and highlight that EPU reduces CO<sub>2</sub>  
288 emissions in the long run. Yu et al. (2021) describe that provincial EPU levels in China have an  
289 adverse impact on CO<sub>2</sub> emissions. Chen et al. (2021), however, document that EPU decreases CO<sub>2</sub>  
290 emissions in both the developed and developing countries, while Abbasi and Adedoyin (2021) use  
291 the dynamic ARDL approach and expound that EPU does not affect CO<sub>2</sub> emissions.

292 Based on the above discussion, it could be noted that the relationship between EPU and the  
293 environment is yet unclear. One of the reasons could be that instead of using disaggregate measures  
294 of EPU (i.e., MU, FU, and TU), the aggregate/composite measure has been extensively employed  
295 for the analysis in the prior literature. Although the study of Alola (2019 a & b) which is the closest  
296 considered the impact of the disaggregate parts of the EPU vis-à-vis the monetary and trade policy  
297 uncertainty on environmental degradation, the model is short of the economic and fiscal policy  
298 uncertainties. However, it might be possible that disaggregated policy uncertainty measures (i.e.,  
299 MU, FU, and TU) have a heterogeneous impact on carbon emissions. Thus, it is inevitable to probe  
300 the disaggregated parts of EPU on carbon emissions.

### 301 **The model**

302 The empirical analysis has recently recommended an EPC framework, which expounds the  
303 presence of a negative relationship between unemployment and environmental quality (Kashem  
304 and Rahman, 2020; Anser et al., 2021a). Furthermore, it also augments EPC with the EKC  
305 framework proposed by Narayan and Narayan (2010). In their study, they present an EKC  
306 framework based on short- and long-run parameters. That is, if the value of the long-run coefficient  
307 of economic growth is lower than its short-run counterpart, then the EKC hypothesis does hold  
308 (Danish et al., 2020). Next, the analysis incorporates MU, FU, and TU into the modelling approach  
309 to examine whether these variables affect CO<sub>2</sub> emissions in the US. It also adds energy  
310 consumption as a control variable and the empirical model yields:

$$311 \quad CO_2 = f(UNE, IPI, ENC, TU, MU, FU) \quad (1)$$

312 where, CO<sub>2</sub>, UNE, IPI, ENC, TU, MU, and FU are CO<sub>2</sub> emissions, the unemployment rate, the  
313 industrial production index (a proxy of economic growth), energy consumption, trade policy  
314 uncertainty, monetary policy uncertainty, and fiscal policy uncertainty, respectively. The  
315 envisaged sign of UNE is expected to be negative, implying that higher unemployment mitigates  
316 the production of goods and services and, hence, impedes CO<sub>2</sub> emissions (Kashem and Rahman,  
317 2020). IPI proxies economic growth, with the expected sign being positive, implying that industrial  
318 production upsurges carbon emissions (Syed and Bouri, 2021). Next, the expected sign of ENC is  
319 positive, which shows that increased consumption of fossil fuels deteriorates the environment by  
320 generating higher CO<sub>2</sub> emissions. The expected signs of TU, FU, and MU are positive, indicating  
321 that policy-related uncertainties escalate carbon emissions (Wang et al., 2020). However, there is  
322 weak empirical evidence for the presence of a negative relationship between policies-related  
323 uncertainty and environmental metrics (Syed and Bouri, 2021).

## 324 The ARDL methodology

325 The complex nature of economic indicators/variables allows them to behave differently in both the  
 326 long- and short-run. That is, there is a likelihood of a positive relationship between variables in the  
 327 long-run, while it is also possible that there exists a negative relationship between the same  
 328 variables in the short-run. To cover this aspect, the ARDL approach has been extensively applied  
 329 in the literature of economics and finance. More specifically, the ARDL model renders both the  
 330 short- and long-run estimates and outperforms other co-integration methodologies for several  
 331 reasons. For instance, the ARDL approach is functional if the variables are co-integrated at  
 332 different orders. Moreover, it explicitly considers the issue of endogeneity, which may cause  
 333 unreliable estimates. The ARDL modelling methodology also renders more reliable results even  
 334 in the case of small samples, while it lets both the dependent and independent variables have  
 335 different optimal lags, thus, providing relatively deep insights for policy implications. The  
 336 equation for the ARDL model yields:

$$\begin{aligned}
 337 \quad \Delta CO_2 = & \alpha + \sum_{i=1}^p \beta_i \Delta CO_{2t-i} + \sum_{i=1}^q \gamma_i \Delta MU_{t-i} + \sum_{i=1}^q \omega_i \Delta FU_{t-i} + \sum_{i=1}^q \psi_i \Delta TU_{t-i} + \\
 338 \quad & \sum_{i=1}^q \delta_i \Delta IPI_{t-i} + \sum_{i=1}^q \varphi_i \Delta UNE_{t-i} + \sum_{i=1}^q \theta_i \Delta ENC_{t-i} + \pi_1 CO_{2t-1} + \pi_2 MU_{t-1} + \pi_3 FU_{t-1} + \\
 339 \quad & \pi_4 IPI_{t-1} + \pi_5 UNE_{t-1} + \pi_6 ENC_{t-1} + \varepsilon_t \quad (2)
 \end{aligned}$$

340 where GDP shows GDP per capita, GDP2 represents the squared GDP, and ENC represents energy  
 341 consumption, respectively.  $\Delta$  denotes first differences,  $i$  represents time lag,  $t$  denotes time, and  $\varepsilon_t$   
 342 is the error term. In addition,  $\pi_i$  ( $i = 1, \dots, 6$ ) shows long-run elasticity, with  $\beta$ ,  $\gamma$ ,  $\omega$ ,  $\psi$ ,  $\delta$ ,  $\varphi$ , and  $\theta$   
 343 being short-run parameters.



344 The envisaged signs of GDP and GDP2 are positive and negative, respectively. This shows  
 345 that EKC does exist (Syed and Bouri, 2021). In addition, the expected coefficient of ENC is  
 346 positive, implying that energy consumption leads to higher CO<sub>2</sub> emissions (Anser et al., 2021b).  
 347 Yet, the expected signs of MU, FU, and TU remain unknown.

### 348 **The dynamic ARDL simulations approach**

349 Although the ARDL approach renders both short- and long-run estimates, the complex dynamic  
 350 nature of ARDL modeling causes a few inconveniences, while explaining and/or interpreting the  
 351 coefficients. Moreover, the inclusion of multiple lags, differences, and lag differences may cause  
 352 complexities in the ARDL approach (Jordan and Philips, 2018). To overcome these issues, Jordan  
 353 and Philips (2018) put forward the dynamic ARDL simulations approach. This method uses  
 354 simulations to render more robust and efficient outcomes. In addition, this novel approach allows  
 355 investigating the effect of positive and negative shocks in the independent variables on the  
 356 dependent variable through dynamic plots. The dynamic ARDL approach often uses 5,000  
 357 simulations, and, hence, the analysis also applies the same number. Furthermore, the analysis  
 358 selects the optimal number lags on the basis of the AIC criterion. Additionally, to check the overall  
 359 performance of the model, the analysis employs several diagnostic tests, i.e., the Breusch-Godfrey  
 360 Lagrange multiplier (LM) test, the Breusch-Pagan-Godfrey (BG) test, and the Jarque-Bera test.  
 361 The error correction form of the dynamic ARDL model yields Equation (3) below:

$$\begin{aligned}
 362 \quad \Delta CO_2 = & \alpha + \sum_{i=1}^p \beta_i \Delta CO_{2t-i} + \sum_{i=1}^q \gamma_i \Delta MU_{t-i} + \sum_{i=1}^q \omega_i \Delta FU_{t-i} + \sum_{i=1}^q \psi_i \Delta TU_{t-i} + \\
 363 \quad & \sum_{i=1}^q \delta_i \Delta IPI_{t-i} + \sum_{i=1}^q \varphi_i \Delta UNE_{t-i} + \sum_{i=1}^q \theta_i \Delta ENC_{t-i} + \pi_1 CO_{2t-1} + \pi_2 MU_{t-1} + \pi_3 FU_{t-1} + \\
 364 \quad & \pi_4 IPI_{t-1} + \pi_5 UNE_{t-1} + \pi_6 ENC_{t-1} + \varepsilon_t \quad (3)
 \end{aligned}$$

365 It is worth mentioning that both Equation (2) and (3) are similar and they are presented in a  
366 standard format; however, the approach to estimate them is quite different.

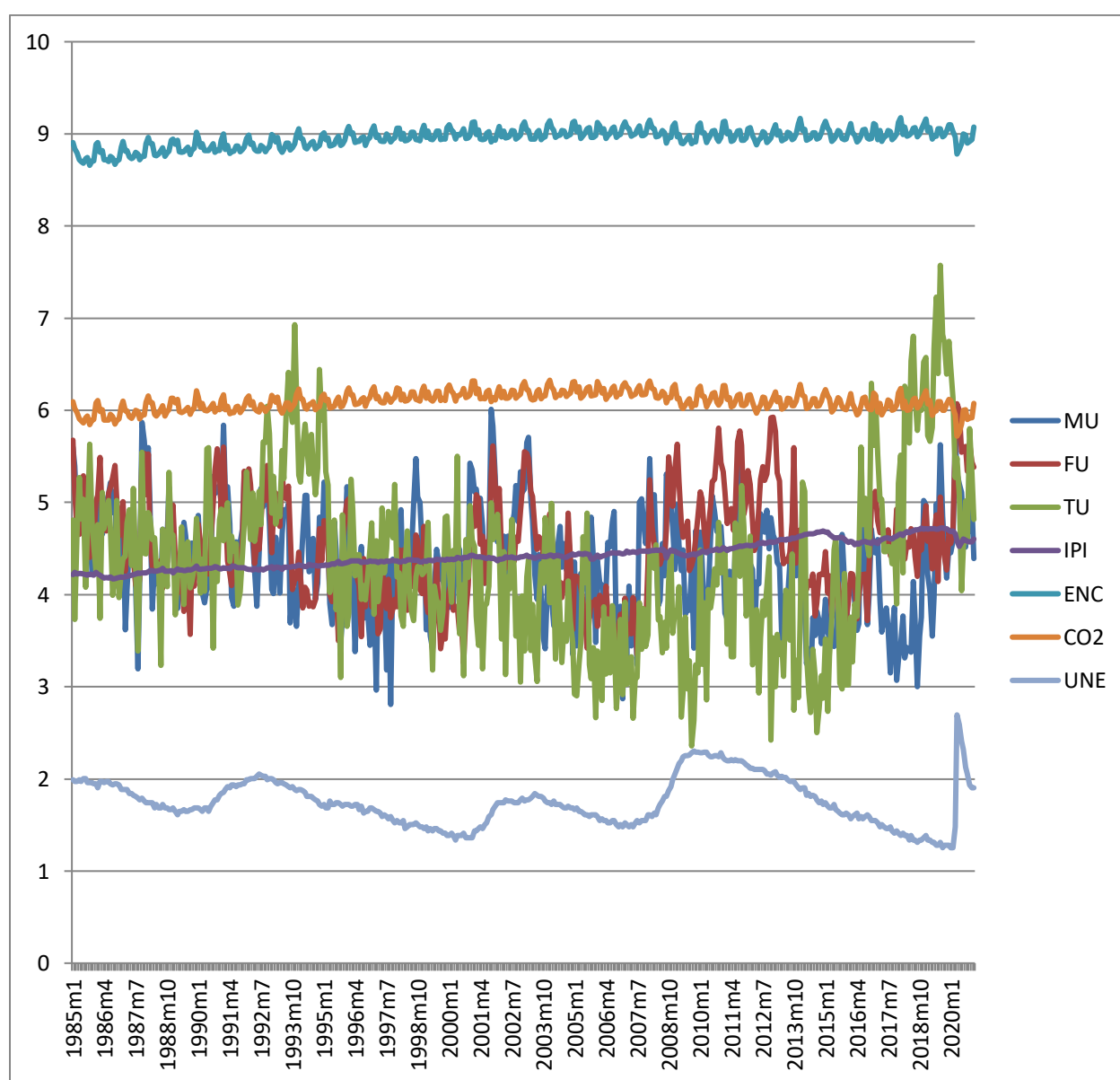
### 367 **Data**

368 The objective of this study is to probe the validity of the EPC hypothesis, along with the role of  
369 MU, FU, and TU in the case of the US. Hence, to this end, the analysis makes use of monthly data,  
370 spanning the period 1985 to 2018. CO<sub>2</sub> emissions are measured in metric tons, while industrial  
371 production is measured with respect to 2012 as the base year, and energy consumption is measured  
372 in oil equivalent per capita. IPI is used as a proxy for GDP, since monthly data of GDP are not  
373 available. Syed and Bouri (2021) also employ IPI as a proxy for GDP and examine the validity of  
374 the EKC hypothesis.

375 Next, the key independent variables are the unemployment rate (% of the labor force), the  
376 monetary policy uncertainty index, the trade policy uncertainty index, and the fiscal policy  
377 uncertainty index. Monetary policy uncertainty is fluctuations and ambiguity about monetary  
378 policy, especially the ambiguity related to interest rates set by the central bank (the Fed). Similarly,  
379 fiscal policy uncertainty is the uncertainty about tax reforms and government expenditures.  
380 Likewise, trade policy uncertainty is the fluctuations and/or uncertainty related to tariff and non-  
381 tariff barriers, the uncertainty about trade agreements, trade remedy laws, and exchange rate  
382 management. Data on the unemployment rate, energy consumption, carbon emissions, and  
383 industrial production are retrieved from the Federal Reserve Economic Database (FRED) and the  
384 Energy Information Administration (EIA). Data on MU, FU, and TU are obtained from  
385 policyuncertainty.com. It is worth mentioning that MU, FU, and TU are calculated on the basis of  
386 methodology developed by Baker et al. (2016). That is, Baker et al. (2016) calculate policy

387 uncertainty about several economic policies through newspaper articles. For instance, MU is  
 388 calculated through the frequency of the number of related words (e.g., monetary policy,  
 389 uncertainty, interest rate uncertainty, among others) in newspaper articles. The same method is  
 390 used to calculate FU and TU. Next, the trend of selected variables of this study is depicted in Figure  
 391 1.

392 Figure 1: Historical trend of selected variables



393

394 As can be seen from Figure 1, there exists relatively high uncertainty in MU, FU, and TU as  
 395 compare to other variables such as IPI and carbon emissions.

396 The summary of data is reported in Table 1, while summary statistics are reported in Table 2.  
 397 All series are expressed in logs to provide a better interpretation of the estimates, as well as to  
 398 avoid heteroscedasticity issues.

399 **Table 1** Data source and variables

Symbol	Indicator	Measurement	Source
CO <sub>2</sub>	Carbon dioxide emissions	Metric ton	EIA
IPI	Industrial production index	Level of production against the base year 2012	FRED
ENC	Energy consumption	Oil equivalent per capita	EIA
UNE	Unemployment rate	Percentage of labor force	FRED
MU	Monetary policy uncertainty index	Frequency of “monetary policy uncertainty” related words in newspaper articles	Policyuncertainty.com
FU	Fiscal policy uncertainty index	Frequency of “fiscal policy uncertainty” related words in newspaper articles	Policyuncertainty.com
TU	Trade policy uncertainty index	Frequency of “trade policy uncertainty” related words in newspaper articles	Policyuncertainty.com

400 EIA is Energy Information Administration, whereas FRED is Federal Reserve Economic Database.

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405 **Table 2** Descriptive statistics

	<b>MU</b>	<b>FU</b>	<b>CO<sub>2</sub></b>	<b>IPI</b>	<b>UNE</b>	<b>TU</b>	<b>ENC</b>
Mean	4.36	4.51	6.09	4.42	1.75	4.35	8.95
Median	4.04	4.21	6.02	4.13	1/20	4.08	8.08
Std. Dev.	0.59	0.56	0.10	0.14	0.25	0.93	0.10
Skewness	0.72	0.01	0.03	0.00	0.00	0.00	0.00
Kurtosis	0.07	0.02	0.59	0.00	0.74	0.31	0.70
Jarque-Bera	(0.19)	(0.00)***	(0.09)*	(0.00)***	(0.00)***	(0.00)***	(0.00)***

406

407 \*, \*\* and \*\*\* shows level of significance at 10%, 5% and 1% respectively. (.) indicates probability value.

408

409 As can be seen from Table 2, TU shows the standard deviation (i.e., 0.93) across the uncertainty  
410 variables, implying that TU is the most volatile series. In contrast, the standard deviation for CO<sub>2</sub>  
411 emissions and ENC is the lowest, implying that both of them do not drastically fluctuate over time.  
412 Additionally, all series are positively skewed and do not contain heavy tails. It is worth reporting  
413 that all series are non-normally distributed except MU, which follows a normal distribution.

414 Next, the Elliot et al. (1996) unit root test has been employed and the findings are reported in  
415 Table 3. The findings clearly document the presence of a unit root in the levels of all variables  
416 under study, while they turn stationary in their first differences, rendering support for the  
417 employment of the ARDL modelling approach.

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420

421 **Table 3** GLS-ADF unit root test

Variables	t-statistics	
	I(0)	I(1)
CO <sub>2</sub>	-1.28	-4.34***
MU	-2.11	-3.14***
FU	-1.35	-4.26***
TU	-2.41	-5.02***
IPI	-2.05	-4.98***
ENC	-0.12	-4.21***
UNE	-1.89	-2.87**

422

423 \*, \*\*, \*\*\* denotes level of significance at 1%, 5%, and 10%, respectively.

424

425 **Empirical results**426 **The Bounds test**

427 First, Table 4 presents the findings from the bounds test. As can be seen, the estimated values from  
 428 both the F- and t-statistics are greater than the upper bounds values, clearly indicating that there  
 429 exists co-integration among the variables under consideration.

430 **Table 4** Bounds test

Test Statistic	Value	I (0)	I (1)
F-statistic	3.64	2.02	3.24
t-statistic	-4.45	-1.95	-4.04

431 Lower and upper bound values are given at 5%.

**432 ARDL results**

433 Next, we report short- and long-run coefficients from the ARDL model in Table 5. The short-run  
434 coefficients (Panel I) of MU, TU, and FU are statistically insignificant, implying that monetary,  
435 fiscal, and trade policy uncertainty do not affect carbon emissions. The findings are in line with  
436 the conclusion provided by Abbasi and Addedoyin (2021) who note that economic policy  
437 uncertainty does not lead to higher carbon emissions. The reason could be the same magnitude of  
438 the channels/effects that affect carbon emissions. For instance, if the magnitude of the investment  
439 effect is almost equal to the consumption effect, the net effect of policy uncertainty on carbon  
440 emissions is zero. Next, the coefficient of UNE is -0.02 and it is statistically significant, indicating  
441 that a 1% increase in the unemployment rate plunges carbon emissions by 0.02%. That is, a 1%  
442 increase in the unemployment rate plunges carbon emissions about by 1588 metric tons. The  
443 reason for this observation is that increases in unemployment plunges production and, hence,  
444 carbon emissions are expected to decrease. Hence, we report the validity of EPC in the short run  
445 for the case of the US. This finding is also backed by the conclusions reached by Anser et al.  
446 (2021a). Moreover, the coefficients of industrial production and ENC turn out to be 0.17 and 0.98,  
447 respectively, indicating that both upsurge carbon emissions.

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453 **Table 5** ARDL estimates

Indicator	Value	Probability Value
<b>Short-run estimates (Panel I)</b>		
$\Delta MU_t$	0.00	0.10
$\Delta FU_t$	-0.00	0.10
$\Delta IPI_t$	0.17	0.00***
$\Delta UNE_t$	-0.02	0.00***
$\Delta TU_t$	0.11	0.20
$\Delta ENC_t$	0.98	0.00***
$\Delta ENC_{t-1}$	-0.26	0.11
<b>Long-run estimates (Panel II)</b>		
$CO_{2t-1}$	-0.45	0.10
$MU_{t-1}$	0.05	0.02***
$FU_{t-1}$	-0.06	0.02***
$TU_{t-1}$	-0.02	0.03**
$IPI_{t-1}$	-0.24	0.00***
$UNE_{t-1}$	-0.01	0.11
$ENC_{t-1}$	0.81	0.00***
<b>Diagnostics (Panel III)</b>		
$R^2$ (Adjusted)		0.79
Ramsey RESET test		(0.57)
LM test		(0.21)
CUSUM test		Stable
CUSUM <sup>2</sup> test		Stable
Jarque-Bera test		(0.27)



ARCH test	(0.79)
ECT	-0.04***

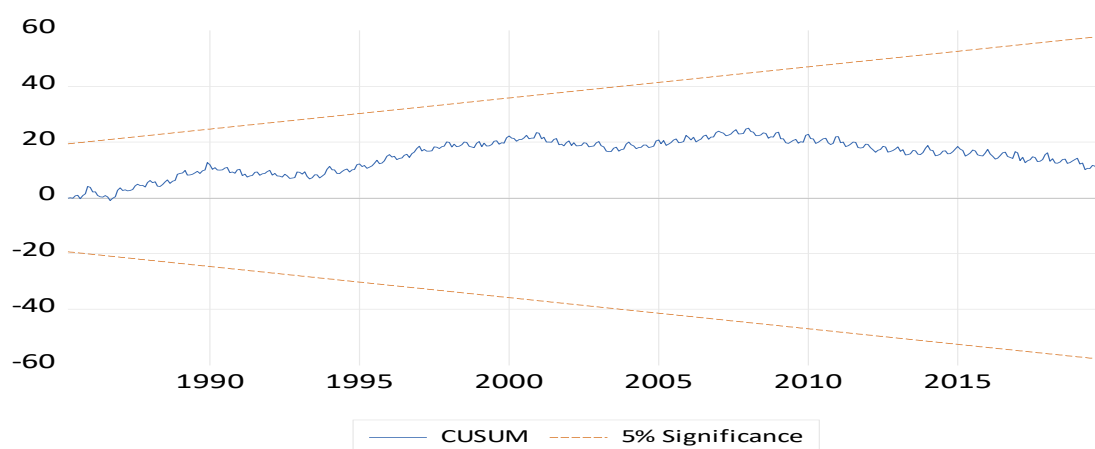
454 \*, \*\*, \*\*\* denotes level of significance at 1%, 5%, and 10%, respectively.

455

456 Regarding the long-run estimates (Panel II), the coefficient of MU is positive and statistically  
457 significant. The value of MU is 0.05, implying that a 1% increase in monetary policy uncertainty  
458 escalates carbon emissions by 0.05%. That is, a 1% increase in MU escalates carbon emissions by  
459 approximately 615 metric tons. This result is in line with those by Danish et al. (2020), who report  
460 that the EPU index escalates carbon emissions. Similarly, the coefficient of FU is negative (i.e., -  
461 0.06) and statistically significant, indicating that a 1% increase in fiscal policy uncertainty plunges  
462 carbon emissions by 0.06%. In particular, a 1% upsurge in FU plunges the carbon emissions by  
463 approximately 738 metric tons. This outcome is backed by the results of Yu et al. (2021), who note  
464 that provincial levels of EPU mitigate carbon emissions. In addition, the coefficient of TU is  
465 negative (i.e., -0.02) and statistically significant, displaying that plunge in carbon emissions by  
466 0.02% is fostered by a 1% increase in trade policy uncertainty. This implies that a 1% increase in  
467 TU impedes carbon emissions by approximately 246 metric tons. This conclusion is in line with  
468 the findings by Chen et al. (2021), with their studies using aggregate measures of economic policy  
469 uncertainty, while ours use disaggregated measures. The long-run coefficient of industrial  
470 production is negative and statistically significant. Since the short- and long-run coefficients of IPI  
471 are respectively positive and negative, we report that EKC does not hold. Furthermore, the  
472 coefficient of ENC is positive and statistically significant, noting that energy consumption leads  
473 to higher carbon emissions.

474 Next, in Panel III (Table 5) we report several diagnostics of the ARDL model for reliability  
 475 and robustness purposes. In order to probe the model's misspecification, the analysis applies the  
 476 Ramsey's RESET test. Its value is 0.57, indicating that the model is well specified. Next, it  
 477 employs the LM serial correlation test to discern whether there exists serial correlation between  
 478 the errors. Its value is 0.21, implying that the errors are not correlated. The CUSUM and CUSUM-  
 479 square tests also explore the stability of the model. Their graphical representations are reported in  
 480 Figures 2.1 and 2.2. As can be seen from both tests, the blue line lie within the critical bounds (red  
 481 lines), thus, we can safely conclude that the exhibits satisfactory stability. The ARCH test, which  
 482 shows heteroscedasticity, notes that the errors variance is constant, indicating that the model does  
 483 not suffer from heteroscedasticity. Finally, the ECT (error correction term) is -0.04, which is  
 484 statistically significant, with its value expounding that any shock to the long-run equilibrium is  
 485 corrected by 4% each month.

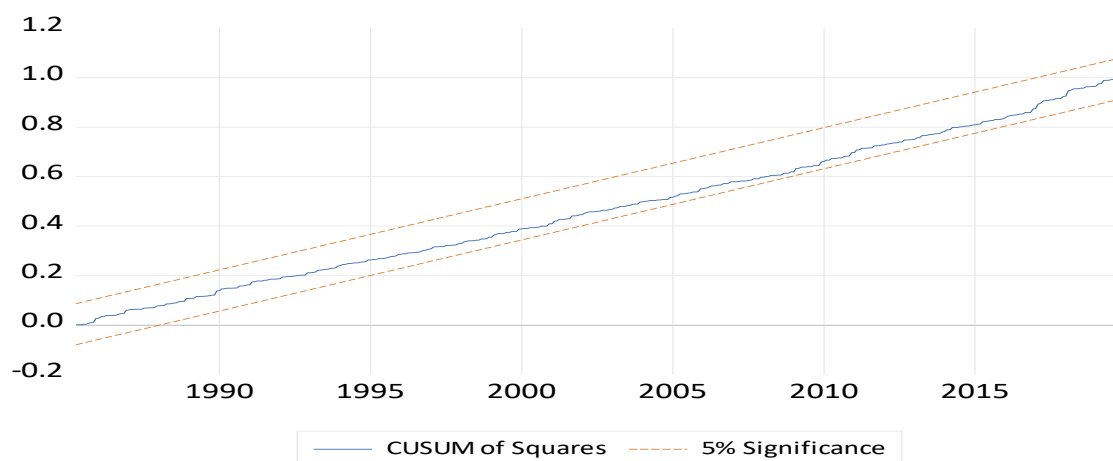
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488 **Fig. 2.1** CUSUM test

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491 **Fig. 2.2** CUSUM-square test

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493 **The dynamic ARDL findings**

494 This section presents the results from the novel dynamic ARDL approach, with the optimal lags

495 being selected through the AIC. The short- and long-run estimates are reported in Table 6. Panel I

496 presents the short-run estimates. More specifically, the coefficient of MU is 0.002 and it is

497 statistically significant, implying that a 1% surge in monetary policy uncertainty escalates carbon

498 emissions by 0.002%. This reports that a 1% increase in MU surges emissions by 246 metric tons.

499 This finding is in line with the conclusions of Amin and Dogan (2021) who highlight that aggregate

500 measures of policy uncertainty (i.e., economic policy uncertainty-EPU) upsurge carbon emissions.

501 The reason might be the fact that monetary policy uncertainty plunges investments in clean energy,

502 R&amp;D, and technology advancements which in turn upsurge carbon emissions. Furthermore, the

503 cumulative magnitude of the investment effect, the innovation channel, the share of the fossil fuel

504 channel, and the energy intensity channel is higher than the combine magnitude of those channels

505 that impedes carbon emissions. In contrast, the coefficient of FU is negative and statistically

506 significant. In particular, the coefficient of FU expounds that a 1% increase in fiscal policy

507 uncertainty plunges emissions by 0.003%. That is, carbon emissions decrease by almost 36 metric

508 tons due to a 1% increase in FU. This outcome is similar to the conclusion of Syed and Bouri  
509 (2021) and Chen et al. (2021). Overall, we can conclude that the magnitude of the consumption  
510 effect is higher than all those channels that escalate emissions. The potential reason for this finding  
511 could be that fiscal policy uncertainty may plummet energy consumption and economic growth  
512 and hence emissions are expected to decrease. Next, the coefficient of TU is statistically  
513 insignificant, indicating that trade policy uncertainty does not affect carbon emissions. This finding  
514 is somehow similar to the results of Abbasi and Addedoyin (2021), who report that policy  
515 uncertainty does not affect emissions. Furthermore, the coefficient of UNE is statistically  
516 insignificant, which indicates that unemployment does not affect the environment, thus providing  
517 evidence that EPC does not hold in the short run. Next, the coefficient of IPI and ENC are negative  
518 and positive, respectively. This highlights that industrial production plunges emissions, while  
519 energy consumption escalates them.

520       Regarding the long-run estimates (Panel II), the coefficients of MU and FU are 0.002 and  
521 -0.002, respectively, indicating that MU escalates carbon emissions, whereas FU plunges them. In  
522 contrast, the coefficient of TU is statistically insignificant, implying that trade policy uncertainty  
523 does not alter these emissions. It is worth mentioning that both the short- and long-run results for  
524 MU, FU, and TU are almost similar. The value of UNE is -0.023 and it is statistically significant,  
525 indicating that a 1% increase in unemployment plunges emissions by 0.023%. That is, a 1%  
526 increase in unemployment rate surges the emissions by 282 metric tons. This outcome  
527 complements the conclusion of Kashem and Rahman (2020) and Anser et al. (2021a). In fact, the  
528 rise in the unemployment rate plunges both production and energy consumption. As a result,  
529 carbon emissions are expected to be decreased. By contrast, both IPI and ENC escalate emissions

530 in long run. Short- and long-run coefficients of IPI are negative and positive, respectively, which  
 531 invalidates the validity of EKC in the case of the US.

532 Related to the diagnostic tests, ECT is -0.07 and statistically significant, highlighting that  
 533 a shock makes carbon emissions to converge by 7% each month. The probability value of the  
 534 model's F-statistic is also statistically significant, while both the CUSUM and CUSUM-square  
 535 tests expound that model is stable. Moreover, we plot the impact of a shock in the independent  
 536 variables on the dependent variable through a dynamic ARDL simulations approach. We set 5,000  
 537 simulations as default and we examine the positive and negative shocks of 10% in the independent  
 538 variables.

539 **Table 6** Dynamic ARDL estimates

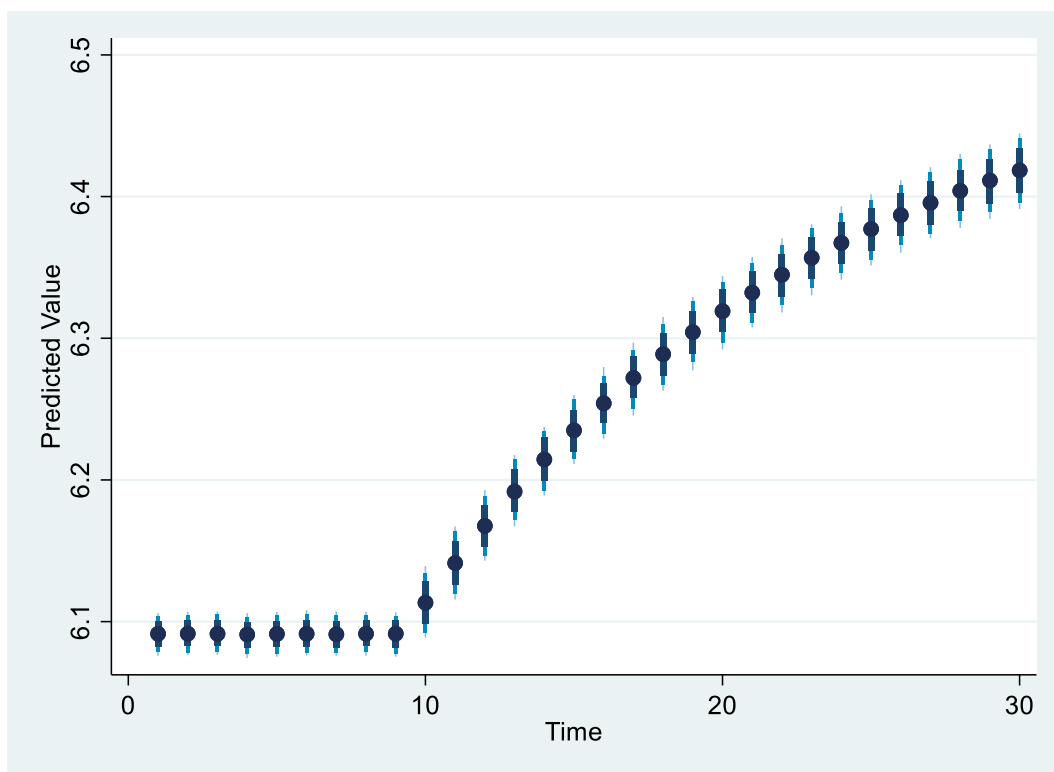
Indicator	Coefficient	p-value
<b>Short-run estimates (Panel I)</b>		
$\Delta MU_t$	0.002	0.00***
$\Delta FU_t$	-0.003	0.00***
$\Delta IPI_t$	-0.021	0.00***
$\Delta UNE_t$	0.009	0.20
$\Delta TU_t$	-0.001	0.16
$\Delta ENC_t$	0.070	0.00***
$\Delta CO_2 t$	-0.260	0.00***
<b>Long-run estimates (Panel II)</b>		
$MU_{t-1}$	0.002	0.02**
$FU_{t-1}$	-0.002	0.02**
$TU_{t-1}$	-0.002	0.13

$IPI_{t-1}$	0.151	0.00***
$UNE_{t-1}$	-0.023	0.00***
$ENC_{t-1}$	1.023	0.00***
<b>Diagnostics (Panel III)</b>		
$R^2$ (Adjusted)		0.80
p-value of F-statistic		(0.00)***
Simulations		5000
CUSUM test		Stable
CUSUM <sup>2</sup> test		Stable
ECT		-0.07***

540 \*, \*\*, \*\*\* denotes level of significance at 1%, 5%, and 10%, respectively. (.) reports p-value

541

542 Figures 3 and 4 illustrate that a shock of +10% in MU escalates carbon emissions in both the short  
543 and long run. In addition, a shock of -10% in MU plunges emissions in both the short and long  
544 run. Furthermore, Figures 5 and 6 expound the shock in FU, and its impact on the predicted value  
545 of carbon emissions. A positive shock in FU impedes emissions in the short, as well as the long  
546 run. In contrast, a negative shock in FU surges emissions in both the short and long run.



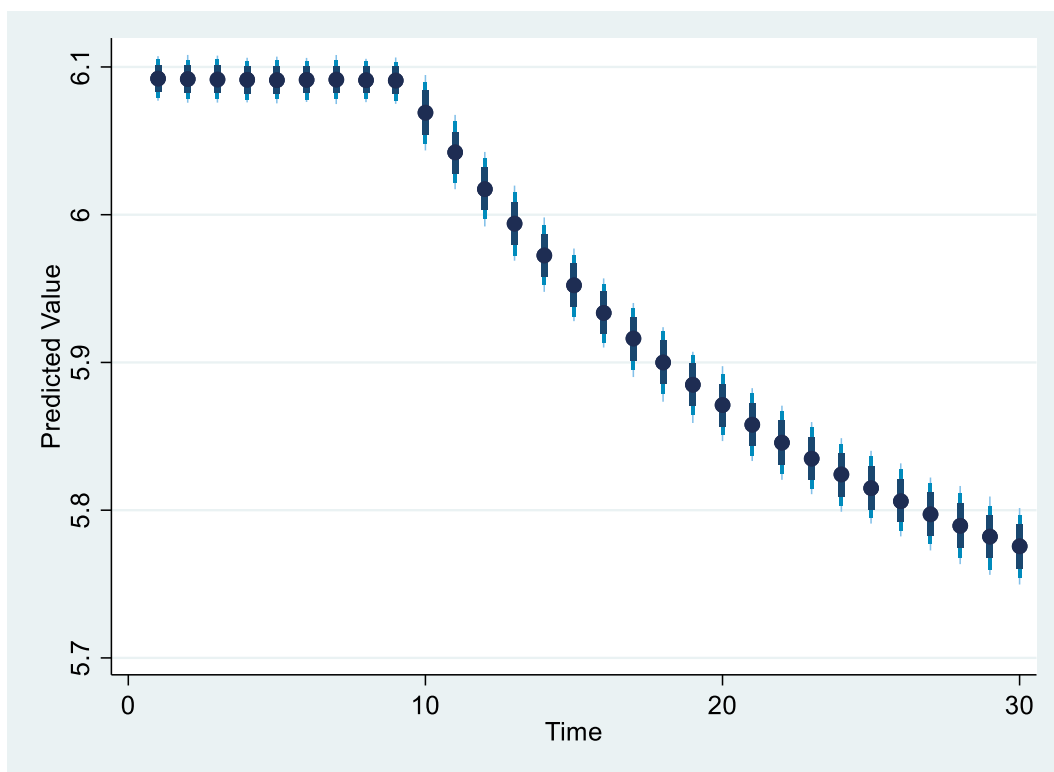
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548 **Fig. 3.** +10% shock in monetary policy uncertainty

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554 **Fig. 4.** A shock of -10% in monetary policy uncertainty

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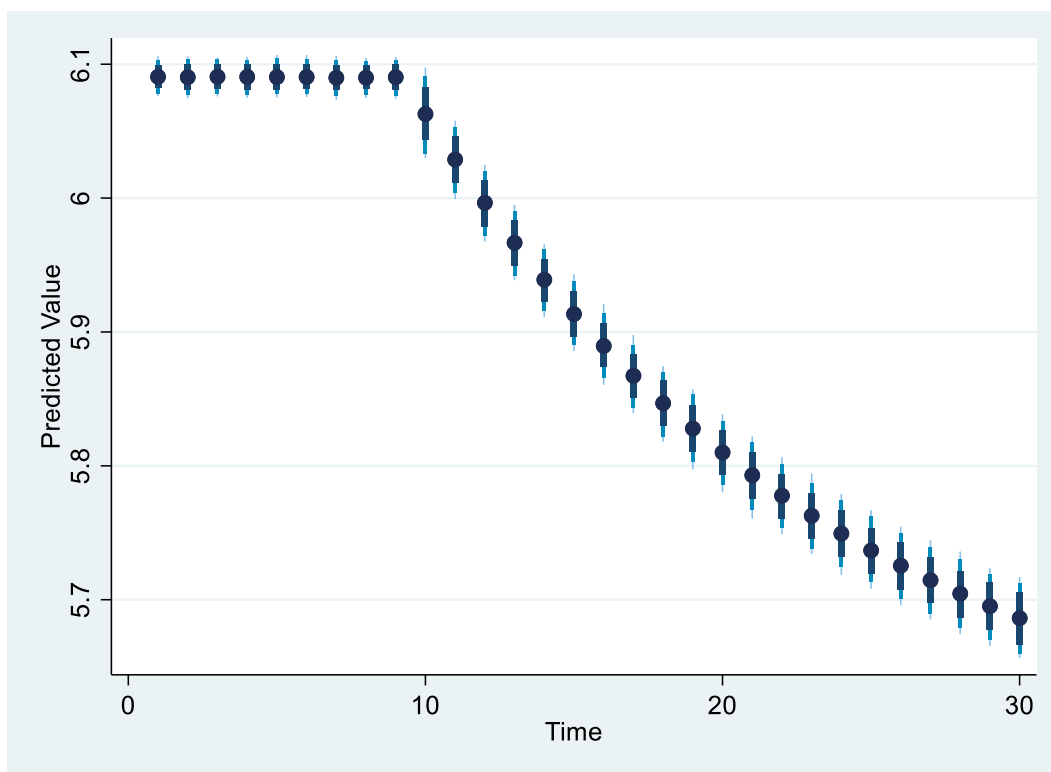
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567 **Fig. 5.** A shock of +10% in fiscal policy uncertainty

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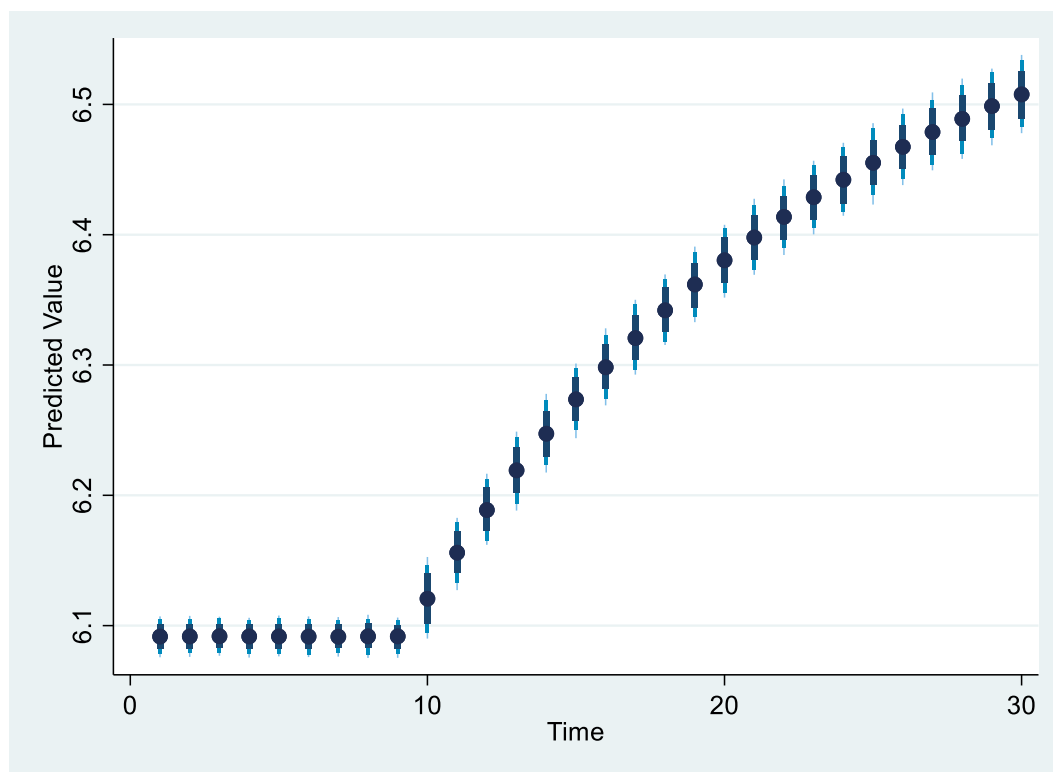
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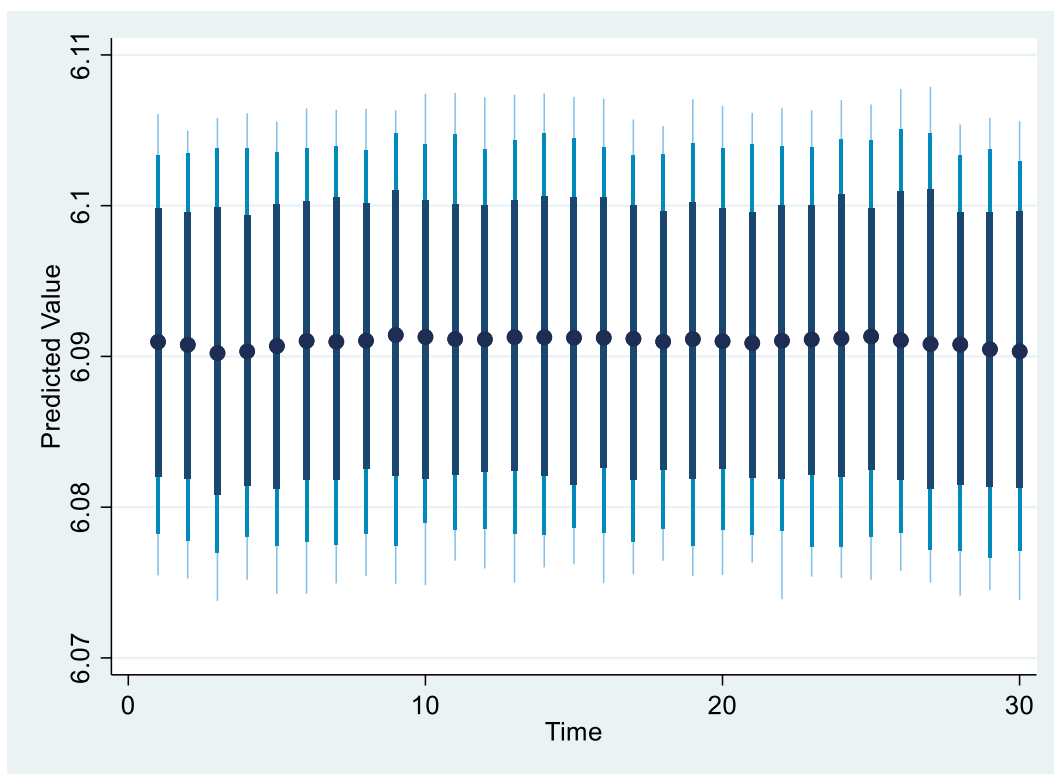


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581 **Fig. 6.** A shock of -10% in fiscal policy uncertainty

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583 Similarly, as can be seen from Figures 7 and 8, a shock in trade policy uncertainty does not affect  
584 the predicted value of carbon emissions, while Figure 9 depicts that a positive 10% shock in  
585 unemployment slightly increases emissions in the short run, while it plunges them in the long run.  
586 In contrast, Figure 10 highlights that a negative shock in unemployment plunges emissions in the  
587 short run, while it escalates them in the long run.



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589 **Fig. 7.** A +10% shock in trade policy uncertainty

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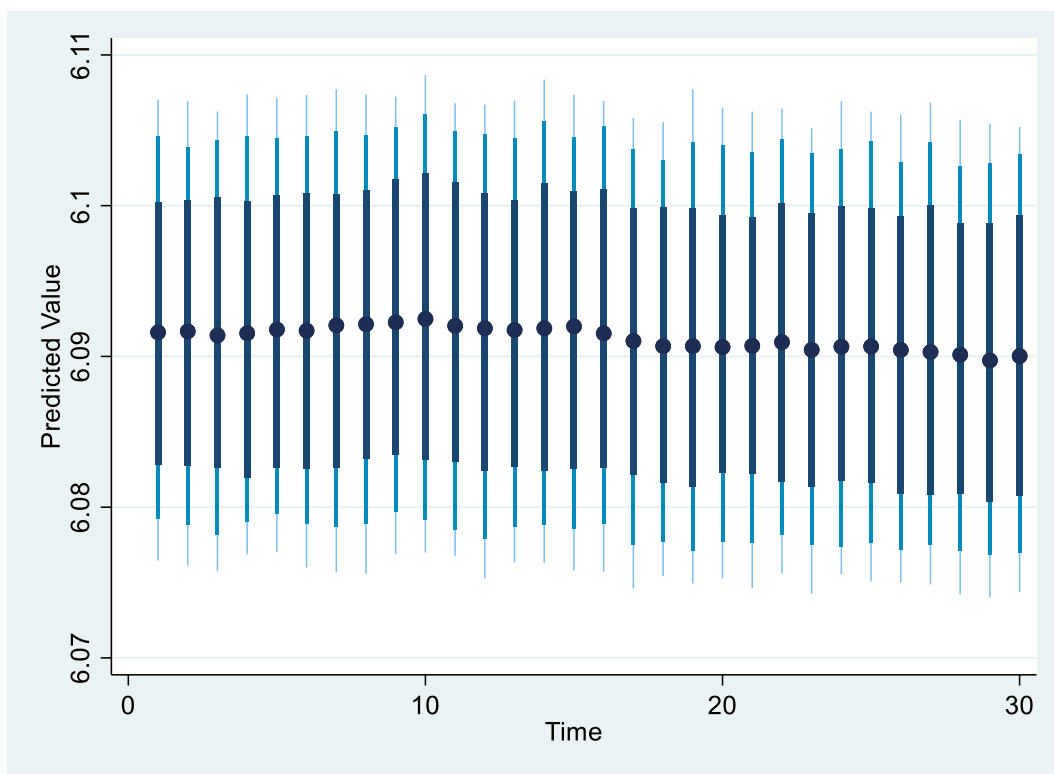
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603 **Fig. 8.** A -10% shock in trade policy uncertainty

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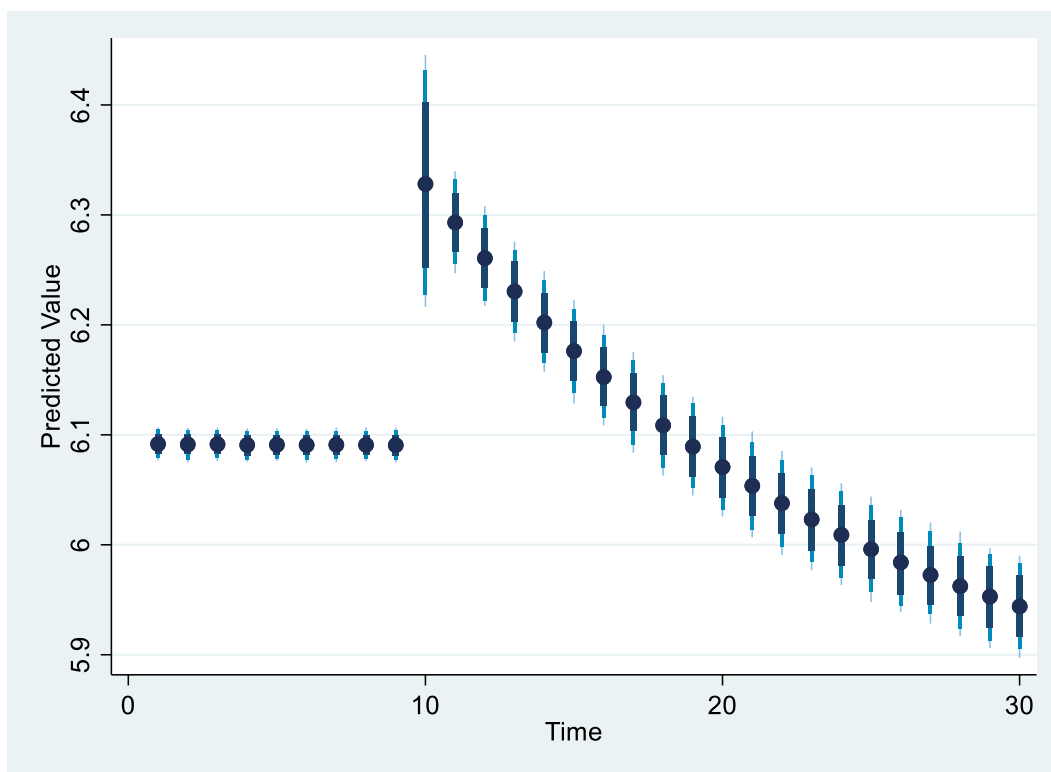
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617 **Fig. 9.** A shock of +10% in unemployment

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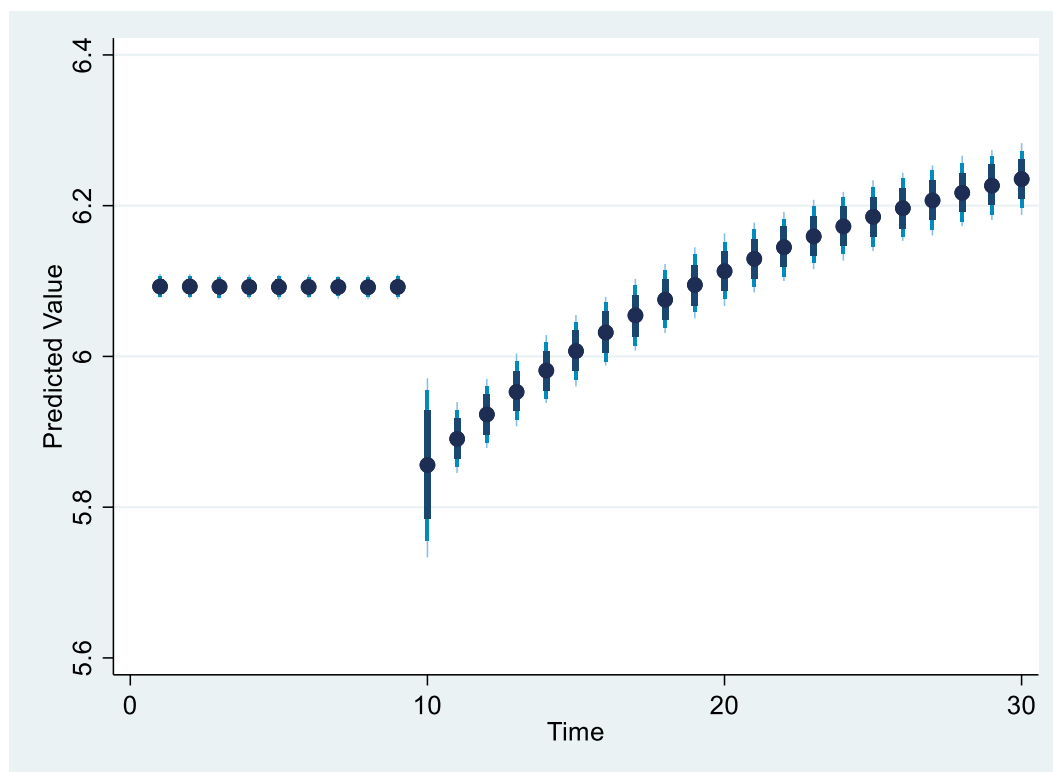
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631 **Fig. 10.** A shock of -10% in unemployment

632 The summary of results is reported in Table 7. As can be seen that EPC is found to be valid in the  
 633 long-run, however, it does not exist in the short-run. Moreover, MU and FU respectively increase  
 634 and decreases carbon emissions in both the short- and long-run. Interestingly, TU affects carbon  
 635 emissions neither in the short- nor in the long-run. Finally, IPI escalates the emissions only in the  
 636 long-run.

637 **Table 7** Summary of Findings from dynamic ARDL simulations

Variable	Short-run	Long-run
IPI	Negative	Positive
ENC	Positive	Positive
UNE	No relationship	Negative

MU	Positive	Positive
FU	Negative	Negative
TU	No relationship	No relationship

638

639 **Conclusion**

640 The goal of this study was to probe the validity of the Environmental Phillips Curve (EPC) in the  
641 US to investigate the impact of monetary policy uncertainty (MU), trade policy uncertainty (TU),  
642 and fiscal policy uncertainty (FU) on carbon dioxide emissions. To this end, the analysis employed  
643 the novel methodology of dynamic ARDL simulations. The findings showed that EPC did not hold  
644 in the short run; however, there its validity was confirmed in the long run. Furthermore, MU  
645 escalated emissions in both the short and long run. In contrast, FU mitigated emissions in the short  
646 run, as well as in the long run. In addition, TU did not affect carbon emissions either in the short  
647 or in the long run. The result of the impact negative impact of FU on environment degradation  
648 could largely be related with the executive decisions vis-à-vis political party partisanship on issues  
649 of energy and climate change. Moreover, the monetary policy which is at the control of the Federal  
650 Reserves are sure determinants of economic productivity, giving the sensitivity of the United  
651 States economy to interest rates and other monetary policy drivers.

652 The study also deduces certain policy implications. Given that decreases in unemployment  
653 lead to escalated emissions, policymakers should further expand innovative policy measures such  
654 as energy efficiency and environmental technologies (e.g., renewable energy, innovations, and  
655 clean technologies that include carbon sequestration and storage) that expectedly spur economic  
656 productivity in order to substantially mitigate the level of emissions without any detrimental effects

657 on unemployment. With such policy, economic stakeholders could be encouraged through  
658 subsidies and tax waiver/reduction on energy-related importations and investments in energy  
659 technologies that are harmless to the environment while also creating sustainable employment and  
660 entrepreneurial opportunities. Concerning the role of various forms of economic uncertainty,  
661 particularly in relevance to those that exerts a negative impact on carbon emissions, the derived  
662 findings also suggest that policymakers (i.e., monetary authorities/central banks) should strive to  
663 maintain the continuity and stability of such desirable policies in order minimize their negative  
664 impact both on carbon emissions and on unemployment. To a large extent, undesirable results are  
665 better avoided through the optimization of sustainable energy consumption structure and the  
666 adoption of emission reduction strategies from various sources. When economies face increasing  
667 economic policy uncertainty, they opt to use cheaper and dirtier fossil fuels, and this leads to  
668 increased carbon emission intensities, along with the associated negative repercussions to the  
669 unemployment rate. Overall, policymakers must pay close attention to the impact of policy  
670 uncertainties on their credibility when formulating and changing their policies.

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682 Author contribution

683 Conceptualization: Roni Bhowmik, Qasim Raza Syed, Andrew Alola

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